

# **CMSC 474, Introduction to Game Theory**

## **Introduction to Probability Theory\***

Mohammad T. Hajiaghayi

University of Maryland

\*: Some slides are adopted from slides by Rong Jin

# Outline

- Basics of probability theory
- Bayes' rule
- Random variable and distributions: Expectation and Variance

# Definition of Probability

- **Experiment**: toss a coin twice
- **Sample space**: possible outcomes of an experiment
  - $S = \{HH, HT, TH, TT\}$
- **Event**: a subset of possible outcomes
  - $A = \{HH\}$ ,  $B = \{HT, TH\}$
- **Probability of an event** : an number assigned to an event  $\Pr(A)$ 
  - Axiom 1:  $0 \leq \Pr(A) \leq 1$
  - Axiom 2:  $\Pr(S) = 1$ ,  $\Pr(\emptyset) = 0$
  - Axiom 3: For two events A and B,  $\Pr(A \cup B) = \Pr(A) + \Pr(B) - \Pr(A \cap B)$
  - Proposition 1:  $\Pr(\sim A) = 1 - \Pr(A)$
  - Proposition 2: For every sequence of disjoint events  $\Pr(\bigcup_i A_i) = \sum_i \Pr(A_i)$

# Joint Probability

- For events  $A$  and  $B$ , **joint probability**  $\Pr(AB)$  (also shown as  $\Pr(A \cap B)$ ) stands for the probability that both events happen.
- Example:  $A=\{HH\}$ ,  $B=\{HT, TH\}$ , what is the joint probability  $\Pr(AB)$ ?

Zero

# Independence

- Two events  $A$  and  $B$  are *independent* in case

$$\Pr(AB) = \Pr(A)\Pr(B)$$

- A set of events  $\{A_i\}$  is *independent* in case

$$\Pr\left(\bigcap_i A_i\right) = \prod_i \Pr(A_i)$$

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- Example: Drug test

	Women	Men
Success	200	1800
Failure	1800	200

$A = \{\text{A patient is a Woman}\}$

$B = \{\text{Drug fails}\}$

Will event  $A$  be independent from event  $B$  ?

$\Pr(A)=0.5, \Pr(B)=0.5, \Pr(AB)=9/20$

# Independence

- Consider the experiment of tossing a coin twice
- Example I:
  - $A = \{HT, HH\}$ ,  $B = \{HT\}$
  - Will event A independent from event B?
- Example II:
  - $A = \{HT\}$ ,  $B = \{TH\}$
  - Will event A independent from event B?
- Disjoint  $\neq$  Independence
- If A is independent from B, B is independent from C, will A be independent from C?

Not necessarily, say  $A=C$

# Conditioning

- If  $A$  and  $B$  are events with  $\Pr(A) > 0$ , the *conditional probability of  $B$  given  $A$*  is

$$\Pr(B | A) = \frac{\Pr(AB)}{\Pr(A)}$$



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	Women	Men
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$\Pr(B|A) = ?$

$\Pr(A|B) = ?$

# Conditioning

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- Example: Drug test

A = {Patient is a Woman}

B = {Drug fails}

	Women	Men
Success	200	1800
Failure	1800	200

$\Pr(B|A) = 18/20$

$\Pr(A|B) = 18/20$

- Given A is independent from B, what is the relationship between  $\Pr(A|B)$  and  $\Pr(A)$ ?

$$\Pr(A|B) = \Pr(A)$$

# Conditional Independence

- Event A and B are *conditionally independent given C* in case

$$\Pr(AB|C) = \Pr(A|C)\Pr(B|C)$$

- A set of events  $\{A_i\}$  is conditionally independent given C in case

$$\Pr(\cap_i A_i | C) = \prod_i \Pr(A_i | C)$$

# Conditional Independence (cont'd)

- Example: There are three events: A, B, C
  - $\Pr(A) = \Pr(B) = \Pr(C) = 1/5$
  - $\Pr(A,C) = \Pr(B,C) = 1/25$ ,  $\Pr(A,B) = 1/10$
  - $\Pr(A,B,C) = 1/125$
  - Whether A, B are independent?  $1/5 * 1/5 \neq 1/10$
  - Whether A, B are conditionally independent given C?  
 $\Pr(A|C) = (1/25)/(1/5) = 1/5$ ,  $\Pr(B|C) = (1/25)/(1/5) = 1/5$   
 $\Pr(AB|C) = (1/125)/(1/5) = 1/25 = \Pr(A|C)\Pr(B|C)$
- A and B are independent  $\neq$  A and B are conditionally independent

# Outline

- Basics of probability theory
- Bayes' rule
- Random variables and distributions: Expectation and Variance

# Bayes' Rule

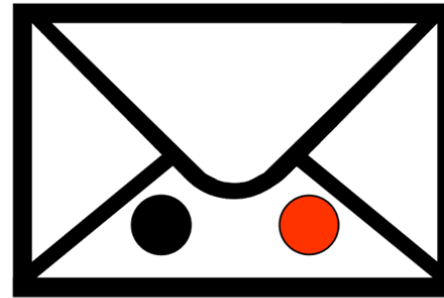
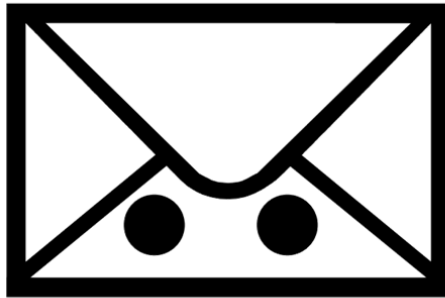
- Given two events A and B and suppose that  $\Pr(A) > 0$ . Then

$$\Pr(A|B) = \frac{\Pr(AB)}{\Pr(B)} = \frac{\Pr(B|A) \Pr(A)}{\Pr(B)}$$

- Why do we make things this complicated?
  - Often  $P(B|A)$ ,  $P(A)$ ,  $P(B)$  are easier to get
  - Some names:
    - **Prior  $P(A)$** : probability before any evidence
    - **Likelihood  $P(B|A)$** : assuming A, how likely is the evidence
    - **Posterior  $P(A|B)$** : conditional prob. after knowing evidence
    - **Inference**: deriving unknown probability from known ones

# Inference with Bayes' Rule: Example

- In a bag there are two envelopes
  - one has a red ball (worth \$100) and a black ball
  - one has two black balls. Black balls worth nothing



- You randomly grabbed an envelope, randomly took out one ball – it's black.
- At this point you're given the option to switch the envelope. **To switch or not to switch?**

# Inference with Bayes' Rule: Example

- E: envelope, 1=(R,B), 2=(B,B)
- B: the event of drawing a black ball
- $P(E|B) = P(B|E)*P(E) / P(B)$
- We want to compare  $P(E=1|B)$  vs.  $P(E=2|B)$
- $P(B|E=1) = 0.5$ ,  $P(B|E=2) = 1$
- $P(E=1)=P(E=2)=0.5$
- $P(B)=3/4$  (it in fact doesn't matter for the comparison)
- $P(E=1|B)=1/3$ ,  $P(E=2|B)=2/3$
- After seeing a black ball, the posterior probability of this envelope being 1 (thus worth \$100) is smaller than it being 2
- Thus you should switch



# Bayes' Rule: More Complicated

- Suppose that  $B_1, B_2, \dots, B_k$  form a partition of  $S$ :

$$B_i \cap B_j = \emptyset; \quad \bigcup_i B_i = S$$

Suppose that  $\Pr(B_i) > 0$  and  $\Pr(A) > 0$ . Then

$$\Pr(B_i | A) = \frac{\Pr(A | B_i) \Pr(B_i)}{\Pr(A)}$$

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$$\begin{aligned} \Pr(B_i | A) &= \frac{\Pr(A | B_i) \Pr(B_i)}{\Pr(A)} \\ &= \frac{\Pr(A | B_i) \Pr(B_i)}{\sum_{j=1}^k \Pr(AB_j)} \end{aligned}$$

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# Outline

- Basics of probability theory
- Bayes' rule
- Random variable and probability distribution: Expectation and Variance

# Random Variable and Distribution

- A *random variable*  $X$  is a numerical outcome of a random experiment
- The *distribution* of a random variable is the collection of possible outcomes along with their probabilities:
  - Discrete case:  $\Pr(X = x) = p_{\theta}(x)$
  - Continuous case:  $\Pr(a \leq X \leq b) = \int_a^b p_{\theta}(x)dx$
- The *support* of a discrete distribution is the set of all  $x$  for which  $\Pr(X=x) > 0$
- The *joint distribution* of two random variables  $X$  and  $Y$  is the collection of possible outcomes along with the joint probability  $\Pr(X=x, Y=y)$ .

# Random Variable: Example

- Let  $S$  be the set of all sequences of three rolls of a die. Let  $X$  be the sum of the number of dots on the three rolls.
- What are the possible values for  $X$ ?
- $\Pr(X = 3) = 1/6 * 1/6 * 1/6 = 1/216,$
- $\Pr(X = 5) = ?$

# Expectation

- A random variable  $X \sim \Pr(X=x)$ . Then, its expectation is

$$E[X] = \sum_x x \Pr(X = x)$$

- In an empirical sample,  $x_1, x_2, \dots, x_N$ ,

$$E[X] = \frac{1}{N} \sum_{i=1}^N x_i$$

- Continuous case:  $E[X] = \int_{-\infty}^{\infty} xp_{\theta}(x)dx$
- In the discrete case, expectation is indeed the average of numbers in the support weighted by their probabilities
- Expectation of sum of random variables

$$E[X_1 + X_2] = E[X_1] + E[X_2]$$

# Expectation: Example

- Let  $S$  be the set of all sequence of three rolls of a die. Let  $X$  be the sum of the number of dots on the three rolls.
- Exercise: What is  $E(X)$ ?
  
- Let  $S$  be the set of all sequence of three rolls of a die. Let  $X$  be the product of the number of dots on the three rolls.
- Exercise: What is  $E(X)$ ?



# Variance

- The variance of a random variable  $X$  is the expectation of  $(X-E[X])^2$  :

$$\begin{aligned}\text{Var}(X) &= E[(X-E[X])^2] \\ &= E[X^2 + E[X]^2 - 2XE[X]] = \\ &= E[X^2] + E[X]^2 - 2E[X]E[X] \\ &= E[X^2] - E[X]^2\end{aligned}$$

# Bernoulli Distribution

- The outcome of an experiment can either be success (i.e., 1) and failure (i.e., 0).
- $\Pr(X=1) = p$ ,  $\Pr(X=0) = 1-p$ , or

$$p_{\theta}(x) = p^x (1-p)^{1-x}$$

- $E[X] = p$ ,  $\text{Var}(X) = E[X^2] - E[X]^2 = p - p^2$

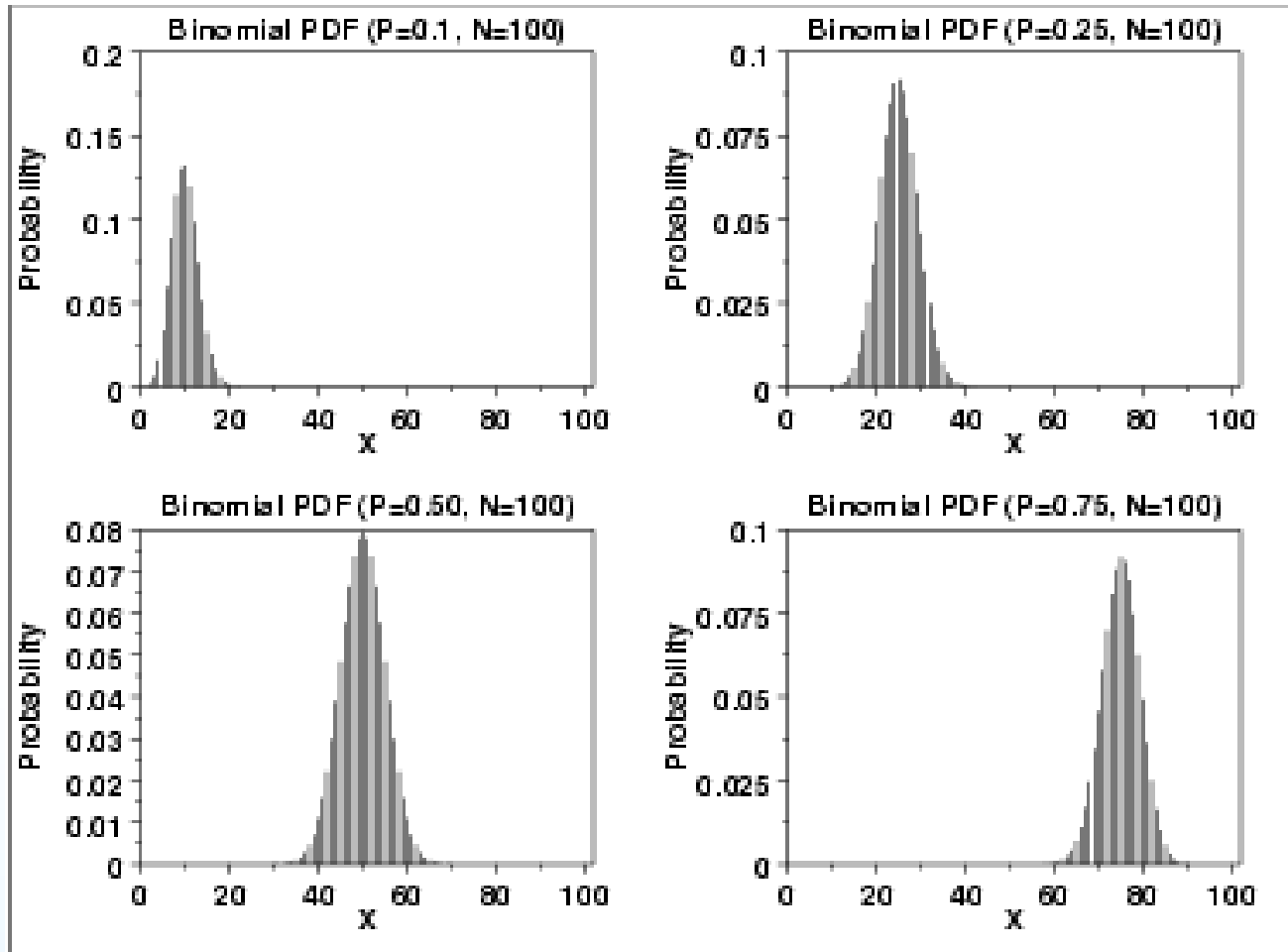
# Binomial Distribution

- $n$  draws of a Bernoulli distribution
  - $X_i \sim \text{Bernoulli}(p)$ ,  $X = \sum_{i=1}^n X_i$ ,  $X \sim \text{Bin}(p, n)$
- Random variable  $X$  stands for the number of times that experiments are successful.

$$\Pr(X = x) = p_{\theta}(x) = \begin{cases} \binom{n}{x} p^x (1-p)^{n-x} & x = 1, 2, \dots, n \\ 0 & \text{otherwise} \end{cases}$$

- $E[X] = np$ ,  $\text{Var}(X) = np(1-p)$

# Plots of Binomial Distribution



# Poisson Distribution

- Coming from Binomial distribution

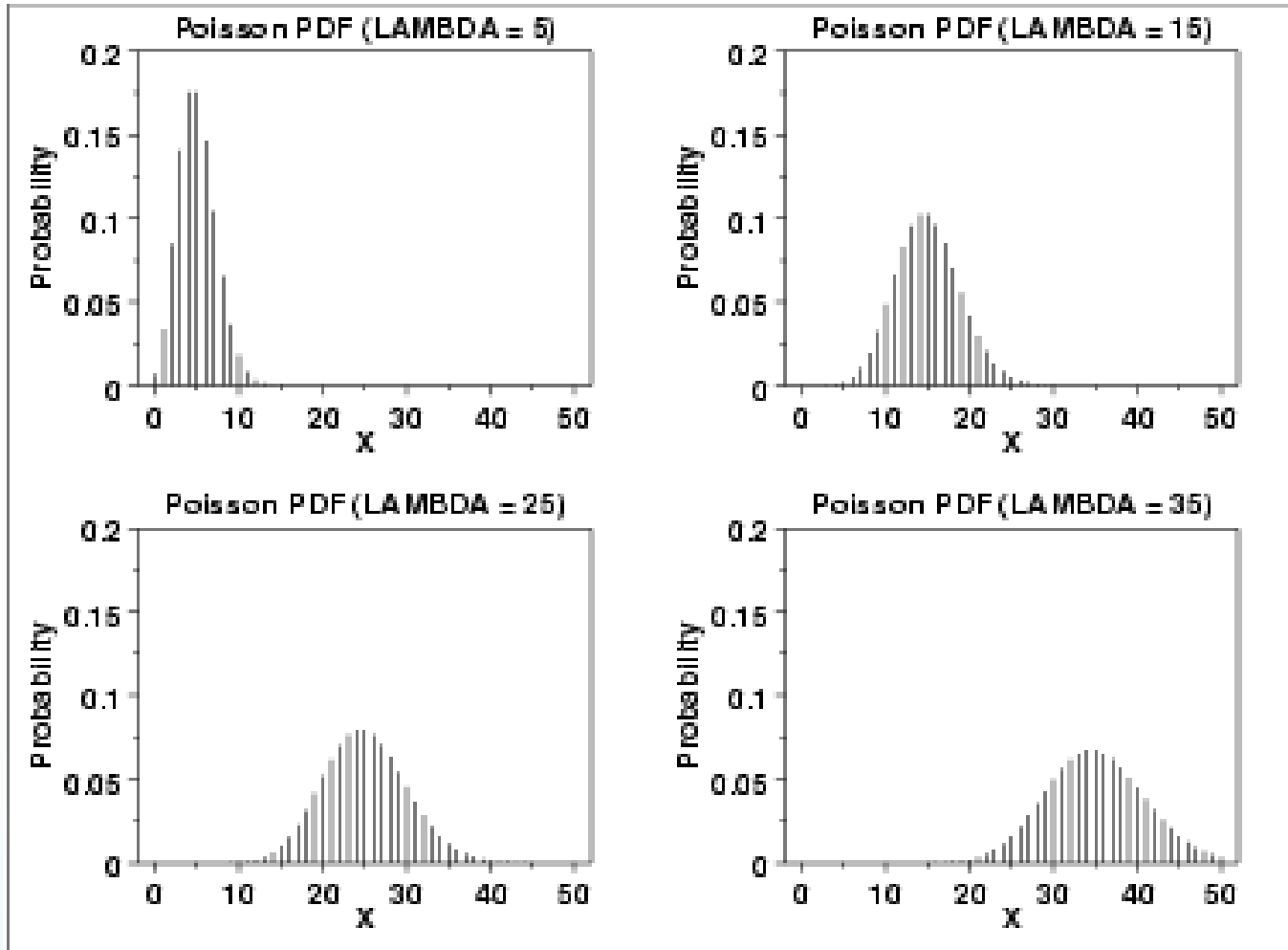
- Fix the expectation  $\lambda=np$
- Let the number of trials  $n \rightarrow \infty$

A Binomial distribution will become a Poisson distribution

$$\Pr(X = x) = p_{\theta}(x) = \begin{cases} \frac{\lambda^x}{x!} e^{-\lambda} & x \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

- $E[X] = \lambda, \text{Var}(X) = \lambda$

# Plots of Poisson Distribution



# Normal (Gaussian) Distribution

- $X \sim N(\mu, \sigma^2)$

$$p_{\theta}(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\}$$

$$\Pr(a \leq X \leq b) = \int_a^b p_{\theta}(x) dx = \int_a^b \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\} dx$$

- $E[X] = \mu$ ,  $\text{Var}(X) = \sigma^2$
- If  $X_1 \sim N(\mu_1, \sigma_1^2)$  and  $X_2 \sim N(\mu_2, \sigma_2^2)$ , for  
 $X = X_1 + X_2 \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$
- Note that Binomial distributions are Normal (Gaussian)